

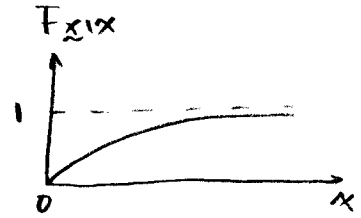
# ECE 4304 HW # 1

$$1. a) \int_0^{\infty} f_X(x) dx = 1 \implies \int_0^{\infty} c \exp(-\alpha x) dx = 1$$

$$\implies c = \alpha$$

$$b) F_X(x) = \int_0^x f_X(x) dx = \int_0^x \alpha \exp(-\alpha x) dx$$

$$= 1 - \exp(-\alpha x) \quad 0 \leq x < \infty$$



when  $\alpha=1$ , the plot is shown on right.

$$c) E[X] = \int_0^{\infty} x f_X(x) dx = \int_0^{\infty} x \alpha \exp(-\alpha x) dx$$

$$= \frac{1}{\alpha}$$

$$d) \text{Var}[X] = \int_0^{\infty} (x - \frac{1}{\alpha})^2 f_X(x) dx$$

$$= \frac{1}{\alpha^2}$$

$$2. P_e = P(0 \text{ sent}) P(\text{got } 1 / 0 \text{ sent})$$

$$+ P(1 \text{ sent}) P(\text{got } 0 / 1 \text{ sent})$$

$$= P(V=-1) P(X > 0 | V=-1) + P(V=1) P(X < 0 | V=1)$$

$$= 0.5 [P(W-1 > 0) + P(1+W < 0)]$$

$$= 0.5 (4.2906e-004 + 4.2906e-004)$$

$$= 4.2906 \times 10^{-4}$$

3. a)

$$\int_{-1}^1 \int_{-1}^1 f_{X,Y}(x,y) dx dy = 1$$

$$\Rightarrow \int_{-1}^1 \int_{-1}^1 c(x-y)^2 dx dy = 1$$

$$\Rightarrow c = \frac{3}{8}$$

b)

$$F_{X,Y}(x,y) = \int_{-1}^x \int_{-1}^y f_{u,v}(u,v) du dv$$

$$= \frac{1}{16} (x+1)(y+1)(2x^2 - 3xy + x + 2y^2 + y + 1)$$

$$c) f_X(x) = \int_{-1}^1 f_{X,Y}(x,y) dy$$

$$= \frac{1}{4} + \frac{3}{4}x^2$$

$$f_Y(y) = \int_{-1}^1 f_{X,Y}(x,y) dx$$

$$= \frac{1}{4} + \frac{3}{4}y^2$$

$$d) P[X > 0, Y > 0] = \int_0^1 \int_0^1 f_{X,Y}(x,y) dx dy$$
$$= \frac{1}{16}$$

$$e) \text{Cov}[X,Y] = E[XY]$$

$$= \int_{-1}^1 \int_{-1}^1 xy f_{X,Y}(x,y) dx dy$$

$$= -\frac{1}{3}$$

(2)

$$(f) \mu_x = \int_{-1}^1 x f_x(x) dx = 0$$

the same,  $\mu_y = 0$

$$\begin{aligned} \text{COV}[XY] &= E[XY] - \mu_x \mu_y \\ &= -\frac{1}{3} \end{aligned}$$

4) a)  $Z(t_1)$  and  $Z(t_2)$  are Gaussian RVs, so their joint probability density function is of the form

$$f(z_1, z_2) = \frac{1}{2\pi\sigma_1\sigma_2\sqrt{1-p^2}} \exp \left\{ -\frac{1}{2(1-p^2)} \left( \frac{(z_1-\eta_1)^2}{\sigma_1^2} - 2p \frac{(z_1-\eta_1)(z_2-\eta_2)}{\sigma_1\sigma_2} + \frac{(z_2-\eta_2)^2}{\sigma_2^2} \right) \right\}$$

$p$  is the correlation coefficient

$$\left. \begin{aligned} \eta_1 &= E[Z(t_1)] = \cos(2\lambda t_1) E[X] + \sin(2\lambda t_1) E[Y] \\ E[X] &= E[Y] = 0 \end{aligned} \right\} \Rightarrow E[Z(t_1)] = 0$$

(3)

Similarly  $E[Z(t_2)] = 0$

$$\text{So } \text{Cov}[Z(t_1)Z(t_2)] = E[Z(t_1)Z(t_2)]$$

$$= \cos(2\lambda t_1) \cos(2\lambda t_2) E[X^2]$$

$$+ [\cos(2\lambda t_1) \sin(2\lambda t_2) + \sin(2\lambda t_1) \cos(2\lambda t_2)] E[XY]$$

$$+ \sin(2\lambda t_1) \sin(2\lambda t_2) E[Y^2]$$

where  $E[XY] = 0$  ( $X, Y$  are independent and  $E[X] = E[Y] = 0$ )

$$E[X^2] = \sigma_X^2 + (E[X])^2 = 1$$

$$\text{similarly } E[Y^2] = 1$$

$$\text{Hence } \text{Cov}[Z(t_1)Z(t_2)] = \cos[2\lambda(t_1 - t_2)]$$

when  $t_1 = t_2$  covariance turns into variance

$$\text{so } \sigma_{Z(t_1)}^2 = \sigma_{Z(t_2)}^2 = 1$$

$$\text{Therefore, } \rho = \frac{\text{Cov}[Z(t_1)Z(t_2)]}{\sigma_{Z(t_1)} \sigma_{Z(t_2)}}$$

$$= \cos[2\lambda(t_1 - t_2)]$$

Hence the joint probability density function is

$$f_{Z(t_1)Z(t_2)}(z_1, z_2) =$$

$$\left. \frac{1}{2\lambda \sin[2\lambda(t_1 - t_2)]} \exp \left\{ \frac{1}{2\sin^2[2\lambda(t_1 - t_2)]} (z_1^2 - 2\cos[2\lambda(t_1 - t_2)]z_1z_2 + z_2^2) \right\} \right\}$$

b) The covariance of  $Z(t_1)$  and  $Z(t_2)$  only depends on the time difference, so  $Z(t)$  is wide sense stationary. Since it is Gaussian it is also strictly stationary.

5 a) by observation,

$$F_{X(t)}(x) = \begin{cases} 0, & x < 0 \\ \frac{1}{2}, & 0 \leq x \leq A \\ 1, & x > A \end{cases}$$

correspondingly,

$$f_{X(t)}(x) = \frac{1}{2} \delta(x) + \frac{1}{2} \delta(x-A)$$

b) By ensemble-averaging,

$$\begin{aligned} E[X(t)] &= \int_{-\infty}^{\infty} x f_{X(t)}(x) dx \\ &= \frac{A}{2} \end{aligned}$$

$$\begin{aligned} R_X(z) &= E[X(t+z)X(t)] \\ &= \int_{-\infty}^{\infty} X(t)X(t+z) f_{Td}(td) dt \end{aligned}$$

Consider only one period:

$$X(t) = \begin{cases} A & t - \frac{T_0}{2} < td < t \\ 0 & t < td < \frac{T_0}{2} + t \end{cases}$$

$$X(t+z) = \begin{cases} A & t - \frac{T_0}{2} + z < td < t+z \\ 0 & t+z < td < \frac{T_0}{2} + t+z \end{cases}$$

Hence  $R_X(z) = \begin{cases} A^2 \int_{t-\frac{T_0}{2}+z}^t \frac{1}{T_0} dt & z > 0 \\ A^2 \int_{t-\frac{T_0}{2}}^{t+z} \frac{1}{T_0} dt & z < 0 \end{cases}$

$$= \frac{A^2}{2} \left( 1 - \frac{2|z|}{T_0} \right) \quad |z| \leq \frac{T_0}{2}$$

$R_X(z)$  is periodic with period  $T_0$ .

c) By time-averaging,

By observation,

$$\mu_x = \frac{A}{2}$$

$$R_x(z, T_0) = \frac{1}{T_0} \int_{-\frac{T_0}{2}}^{\frac{T_0}{2}} X(t+z) X(t) dt$$

$$= \frac{A^2}{2} \left( 1 - 2 \frac{|z|}{T_0} \right), \quad |z| \leq \frac{T_0}{2}$$

$R_x(z)$  is periodic with period  $T_0$

d) For the reason that ensemble-averaging and time-averaging yield the same results for the mean and autocorrelation,  $X(t)$  is ergodic in both mean and autocorrelation function. The ergodicity also implies that  $X(t)$  is wide sense stationary.

```
%-----  
%ECE4304 Hw#1  
%problem 6  
%-----  
N=1e6;           % # of iterations  
W=0.3*randn(N,1); %Gaussian Noise with mean=0, standard deviation=0.3  
temp=rand(N,1);  
V=(temp>0.5)-(temp<0.5); %transmit signal  
X=V+W;          %corrupted signal  
Y=sign(X);      %reciever decision  
Pe=1-sum(Y==V)/N %error probability
```