

ECE531 Screencast 12.5: Detection of a Known Signal with Unknown Parameters in a Linear Model

D. Richard Brown III

Worcester Polytechnic Institute

Linear Model

Suppose we have observations given by

$$y = Hx + w$$

where $H \in \mathbb{R}^{n \times p}$ is known, $x \in \mathbb{R}^{p \times 1}$ is a vector of parameters/states, at least some of which are unknown, and $W \in \mathbb{R}^{n \times 1}$ with $W \sim \mathcal{N}(0, \sigma^2 I)$ is AWGN with σ^2 known.

We want to decide between

$$\mathcal{H}_0 : x = 0$$

$$\mathcal{H}_1 : x \neq 0$$

Many problems fit into this linear model, even some that aren't obvious.

Example: Sinusoidal Signal with Unknown Amp. and Phase

Suppose the signal we are trying to detect is

$$s_k = a \cos(\omega k + \phi)$$

for $k = 0, \dots, n - 1$ with unknown a and ϕ . We can write

$$s_k = \alpha_1 \cos(\omega k) + \alpha_2 \sin(\omega k)$$

with $\alpha_1^2 + \alpha_2^2 = a^2$. Letting the unknown parameters $x = [\alpha_1, \alpha_2]^T$, we have

$$y = Hx + w$$

with

$$H = \begin{bmatrix} 1 & 0 \\ \cos(\omega_0) & \sin(\omega_0) \\ \vdots & \vdots \\ \cos(\omega_0(n-1)) & \sin(\omega_0(n-1)) \end{bmatrix}$$

GLRT for Linear Model

Since the MLE for the linear model with AWGN is simply

$$\hat{x} = (H^T H)^{-1} H^T y$$

the GLRT decision rule can be written as follows. We decide \mathcal{H}_1 if

$$\begin{aligned} & \frac{p_1(y; \hat{x})}{p_0(y)} > v \\ \Leftrightarrow & \frac{\exp\left(-\frac{(y-H\hat{x})^T(y-H\hat{x})}{2\sigma^2}\right)}{\exp\left(-\frac{y^T y}{2\sigma^2}\right)} > v \\ \Leftrightarrow & 2\hat{x}^T H^T y - \hat{x} H^T H \hat{x} > v' \\ \Leftrightarrow & 2y^T H (H^T H)^{-1} H^T y - y^T H (H^T H)^{-1} H^T H (H^T H)^{-1} H^T y > v' \\ \Leftrightarrow & y^T H (H^T H)^{-1} H^T y > v' \end{aligned}$$

with v' chosen to satisfy $P_{\text{fp}} \leq \alpha$. The decision statistic is χ^2 distributed, so we can get exact expressions for P_{fp} and P_D .

Bayesian Detection for Linear Model with Gaussian Prior

In the Bayesian case with $Y = Hx + W$, we assume a Gaussian prior on x such that

$$\pi(x) = \pi_0 \delta(x) + (1 - \pi_0) \frac{1}{(2\pi)^{n/2} |\Sigma_x|^{1/2}} \exp\left(-\frac{(x - \mu_x)^\top \Sigma_x^{-1} (x - \mu_x)}{2}\right).$$

If we let $s = Hx$ then s is also Gaussian with $\mu_s = H\mu_x$ and $\Sigma_s = H\Sigma_x H^\top$. We assume μ_s and Σ_s are known and s is independent of the noise $W \sim \mathcal{N}(0, \sigma^2 I)$.

We want to decide between

$$\mathcal{H}_0 : x = 0 \Leftrightarrow s = 0$$

$$\mathcal{H}_1 : x \neq 0 \Leftrightarrow s \neq 0$$

which is equivalent to

$$\mathcal{H}_0 : Y \sim \mathcal{N}(0, \sigma^2 I)$$

$$\mathcal{H}_1 : Y \sim \mathcal{N}(\mu_s, \Sigma_s + \sigma^2 I)$$

This is actually a simple binary hypothesis testing problem!

Bayesian Detection for Linear Model with Gaussian Prior

Since this problem is simple and binary, we know we decide \mathcal{H}_1 when

$$\begin{aligned} \frac{p_1(y)}{p_0(y)} &> \frac{\pi_0}{1 - \pi_0} \\ \Leftrightarrow \frac{\frac{1}{(2\pi)^{n/2} |\Sigma_s + \sigma^2 I|^{1/2}} \exp\left(-\frac{(y - \mu_s)^\top (\Sigma_s + \sigma^2 I)^{-1} (y - \mu_s)}{2}\right)}{\frac{1}{(2\pi)^{n/2} |\sigma^2 I|^{1/2}} \exp\left(-\frac{y^\top y}{2\sigma^2}\right)} &> \frac{\pi_0}{1 - \pi_0} \\ \Leftrightarrow \frac{|\sigma^2 I|^{1/2}}{|\Sigma_s + \sigma^2 I|^{1/2}} \exp\left(-\frac{(y - \mu_s)^\top (\Sigma_s + \sigma^2 I)^{-1} (y - \mu_s)}{2} + \frac{y^\top y}{2\sigma^2}\right) &> \frac{\pi_0}{1 - \pi_0} \\ \Leftrightarrow \frac{y^\top y}{\sigma^2} - (y - \mu_s)^\top (\Sigma_s + \sigma^2 I)^{-1} (y - \mu_s) &> 2 \ln \left(\frac{\pi_0 |\Sigma_s + \sigma^2 I|^{1/2}}{(1 - \pi_0) |\sigma^2 I|^{1/2}} \right) \\ \Leftrightarrow y^\top A y + \mu_s^\top B y &> 2 \ln \left(\frac{\pi_0 |\Sigma_s + \sigma^2 I|^{1/2}}{(1 - \pi_0) |\sigma^2 I|^{1/2}} \right) + \mu_s^\top (\Sigma_s + \sigma^2 I)^{-1} \mu_s \end{aligned}$$

with

$$A = \frac{1}{\sigma^2} I - (\Sigma_s + \sigma^2 I)^{-1} \quad \text{and} \quad B = 2(\Sigma_s + \sigma^2 I)^{-1}$$